WEBVTT

- 1 00:00:00.690 --> 00:00:03.090 <v Laura>All right, let's get started.</v>
- 2 00:00:03.090 --> 00:00:05.100 Thank you, everyone, for coming.
- $3\ 00:00:05.100 \longrightarrow 00:00:07.473$ So let me introduce our speaker today.
- 4 00:00:08.460 --> 00:00:11.430 Ariel Chao is a PhD student in the department
- 5 00:00:11.430 --> 00:00:16.320 of biostatistics, advised by me and Donna Spiegelman.
- $6\ 00:00:16.320 --> 00:00:19.653$ So let me say few things about her, about Ariel.
- 7 00:00:20.670 --> 00:00:23.850 So I've been working with Ariel for three years now
- $8\ 00:00:23.850 \longrightarrow 00:00:27.060$ and I have to say it's been a real pleasure.
- 9 00:00:27.060 --> 00:00:30.300 Ariel is an extraordinary student, very patient,
- $10\ 00:00:30.300 \longrightarrow 00:00:34.020$ definitely her characteristic and independent.
- $11\ 00:00:34.020 \longrightarrow 00:00:36.780$ I've always been impressed by her creativity
- $12\ 00:00:36.780 \dashrightarrow 00:00:40.710$ and the way she would always find solutions by herself.
- $13\ 00:00:40.710 \longrightarrow 00:00:42.240$ We have been having issues
- $14~00{:}00{:}42.240 \dashrightarrow 00{:}00{:}44.550$ with getting data from our collaborators
- $15\ 00:00:44.550 \longrightarrow 00:00:48.240$ and she never gave up and found ways to keep working
- $16\ 00:00:48.240 --> 00:00:50.850$ on what she had while waiting.
- $17\ 00:00:50.850 \longrightarrow 00:00:53.400$ So she deeply cares about the applications
- $18\ 00:00:53.400$ --> 00:00:57.450 and she's working on and she has a great intuition.
- $19\ 00:00:57.450 \longrightarrow 00:00:58.860$ I also been impressed
- $20\ 00{:}00{:}58.860 \dashrightarrow 00{:}01{:}03.060$ on how she can work in several things at the same time.
- 21 00:01:03.060 --> 00:01:05.970 And as you will see today, she's very talented
- $22\ 00:01:05.970 \longrightarrow 00:01:09.565$ and I wish her the best for her future career.
- $23\ 00:01:09.565 \longrightarrow 00:01:12.660$ Before that, today, she will present her work
- 24 00:01:12.660 --> 00:01:14.807 on addressing bias in causal effects,
- $25\ 00:01:14.807 \longrightarrow 00:01:18.480$ estimated underspecified interference sets
- $26\ 00:01:18.480 \longrightarrow 00:01:22.470$ with application to HIV prevention trials.

- 27 00:01:22.470 --> 00:01:25.710 So let's give Ariel a more welcome.
- 28 00:01:25.710 --> 00:01:26.860 Ariel, (crackling drowns out speaker).
- $29\ 00:01:30.390 \longrightarrow 00:01:32.070 < v \text{ Speaker>Let me just add,} < / v >$
- $30\ 00:01:32.070 \longrightarrow 00:01:34.350$ Ariel has a lot of material to present
- 31~00:01:34.350 --> 00:01:38.460 so we decided to not take questions while she's talking
- $32\ 00:01:38.460 \longrightarrow 00:01:40.451$ or she'll never get through the necessarily.
- $33\ 00:01:40.451 \longrightarrow 00:01:41.760$ And then we're gonna allow
- $34\ 00:01:41.760 --> 00:01:44.220$ for around 10 minutes at the end for questions.
- 35 00:01:44.220 --> 00:01:45.360 So write down the questions
- $36\ 00:01:45.360 \longrightarrow 00:01:48.180$ and then we'll try to give as many people a chance
- $37\ 00:01:48.180 \longrightarrow 00:01:49.780$ to ask her questions at the end.
- $38\ 00:01:51.330 --> 00:01:53.100 < v \ Laura > I \ will \ keep \ track \ of \ it. </v>$
- 39 00:01:53.100 --> 00:01:55.380 <v Speaker>I'm just monitoring the chat.</v>
- 40~00:01:55.380 --> 00:01:57.900 < v->Oh yes, can someone, 'cause I don't think I can see you.</v>
- 41 00:01:57.900 --> 00:02:00.363 <v Speaker>I can see you.</v>
- 42 00:02:01.470 --> 00:02:02.303 <v -> All right.</v>
- $43\ 00:02:02.303 \dashrightarrow 00:02:04.830$ So thank you, Laura, and it's been a real pleasure
- $44\ 00:02:04.830 \longrightarrow 00:02:06.390$ working with you as well.
- $45~00:02:06.390 \longrightarrow 00:02:10.080$ So today, I'll be presenting on my dissertation research,
- $46\ 00:02:10.080 \longrightarrow 00:02:13.530$ which is on addressing bias in causal effects
- 47 00:02:13.530 --> 00:02:16.380 estimated under misspecified interference sets.
- $48\ 00{:}02{:}16.380 {\: \hbox{--}\!>\:} 00{:}02{:}18.540$ And we've applied our methods through the analysis
- 49 00:02:18.540 --> 00:02:20.373 of HIV prevention trials.
- $50\ 00:02:22.560 --> 00:02:25.740$ So as an introduction, so interference
- $51\ 00:02:25.740 \longrightarrow 00:02:28.860$ or spillover is often present in either randomized
- $52\ 00:02:28.860 \longrightarrow 00:02:30.510$ or observational studies.

- $53\ 00:02:30.510 \longrightarrow 00:02:32.280$ Whereby interference, we mean
- $54\ 00:02:32.280 \longrightarrow 00:02:34.830$ that a participant's outcome can be determined
- 55 00:02:34.830 --> 00:02:36.570 by not only their own exposure
- $56\ 00:02:36.570 \longrightarrow 00:02:38.370$ but also the exposure of others.
- $57\ 00:02:38.370 \longrightarrow 00:02:41.550$ So a common example is with vaccines.
- 58 00:02:41.550 --> 00:02:44.070 So say, my disease status is not only affected
- 59 00:02:44.070 --> 00:02:45.870 by my own vaccination status,
- $60\ 00:02:45.870 \longrightarrow 00:02:48.870$ but also the vaccination status of others around me.
- 61 00:02:48.870 --> 00:02:51.630 And in the context of HIV prevention trials,
- $62\ 00:02:51.630 \longrightarrow 00:02:54.180$ it's been found in several network-based studies
- $63\ 00:02:54.180 \longrightarrow 00:02:57.000$ that when only some participants of a network
- $64\ 00:02:57.000 \longrightarrow 00:03:00.000$ are trained on say HIV knowledge
- $65\ 00:03:00.000 \longrightarrow 00:03:02.520$ or safe practices, that the members
- $66\ 00:03:02.520 \longrightarrow 00:03:05.130$ who are weren't trained in the network
- 67 00:03:05.130 --> 00:03:07.440 also demonstrated increased knowledge
- $68\ 00:03:07.440 --> 00:03:09.493$ and reduced risk behaviors.
- $69\ 00:03:09.493 \longrightarrow 00:03:12.000$ And this is known as disability effect.
- $70~00:03:12.000 \dashrightarrow 00:03:16.770$ So causal inference, that is conducted the presence
- $71\ 00:03:16.770 \longrightarrow 00:03:19.230$ of interference is often done under assumptions
- $72\ 00:03:19.230 \longrightarrow 00:03:22.380$ on the extent and mechanism of interference.
- $73\ 00:03:22.380 \longrightarrow 00:03:25.200$ And typically, this will require a specification
- $74\ 00:03:25.200 \longrightarrow 00:03:28.110$ of an interference set for each participant.
- $75\ 00:03:28.110 \longrightarrow 00:03:29.370$ Whereby interference sets,
- 76 00:03:29.370 \rightarrow 00:03:32.490 we mean that a group of individuals
- $77\ 00:03:32.490 \longrightarrow 00:03:35.490$ who can affect the outcome of that participant.
- $78\ 00:03:35.490 \longrightarrow 00:03:37.050$ And then to this interference set,
- $79~00{:}03{:}37.050 \dashrightarrow 00{:}03{:}40.620$ we also typically apply an exposure mapping function
- $80\ 00:03:40.620 \longrightarrow 00:03:42.990$ that will take the exposure vector
- 81 00:03:42.990 --> 00:03:44.940 observed in this interference set

- $82\ 00:03:44.940 \longrightarrow 00:03:46.950$ and map it to some scaler quantity.
- $83\ 00:03:46.950 \longrightarrow 00:03:49.203$ And we'll see some examples of this later.
- $84\ 00:03:50.940 \longrightarrow 00:03:53.310$ So existing literature interference sets
- $85\ 00:03:53.310 \longrightarrow 00:03:55.980$ are typically assumed to be correctly specified
- $86\ 00:03:55.980 \longrightarrow 00:03:57.180$ so that the exposures
- $87\ 00:03:57.180 \longrightarrow 00:03:59.130$ that are mapped from these interference sets
- $88\ 00:03:59.130 --> 00:04:01.290$ are also correctly measured.
- 89 00:04:01.290 --> 00:04:04.290 But often, this correctly specifying
- $90\ 00:04:04.290 \longrightarrow 00:04:06.240$ an interference set is challenging.
- 91 00:04:06.240 --> 00:04:09.540 For example, networks can be mismeasured
- 92 00:04:09.540 --> 00:04:12.480 and when interference sets are misspecified,
- $93\ 00:04:12.480$ --> 00:04:15.210 we show under various settings that causal effects estimated
- 94 00:04:15.210 --> 00:04:17.880 by usual purchase are typically biased.
- $95~00:04:17.880 \longrightarrow 00:04:20.010$ And there have been several publications
- 96 00:04:20.010 --> 00:04:22.018 that have addressed this issue.
- $97\ 00:04:22.018 --> 00:04:24.540$ And the majority of these publications aim
- 98 00:04:24.540 --> 00:04:26.910 to first estimate the true networks
- $99\ 00:04:26.910 --> 00:04:28.770$ and then using these estimated networks
- $100\ 00:04:28.770 \longrightarrow 00:04:30.750$ to estimate the causal effects.
- $101\ 00:04:30.750 \longrightarrow 00:04:32.370$ And there have also been methods proposed
- $102\ 00:04:32.370 \longrightarrow 00:04:34.263$ for a sensitivity analysis as well.
- $103\ 00:04:36.090 --> 00:04:38.730$ However, we pursue a different approach where we assume
- $104\ 00:04:38.730 \longrightarrow 00:04:41.130$ that we have a validation study
- $105\ 00{:}04{:}41.130 \dashrightarrow 00{:}04{:}44.040$ in which the true interference sets are measured alongside
- $106\ 00:04:44.040 \longrightarrow 00:04:46.380$ the observed or surrogate ones for a subset
- $107\ 00:04:46.380 \longrightarrow 00:04:47.760$ of the study sample.
- 108 00:04:47.760 --> 00:04:48.810 And this will allow us
- $109\ 00{:}04{:}48.810 \dashrightarrow 00{:}04{:}51.510$ to empirically estimate the measurement error process

- $110\ 00:04:51.510$ --> 00:04:54.540 and use the estimated measurement error parameters
- $111\ 00:04:54.540 \longrightarrow 00:04:57.690$ to bias correct causal effects.
- $112\ 00{:}04{:}57.690 \dashrightarrow 00{:}05{:}00.690$ So again, this dissertation is a collection of three papers
- $113\ 00:05:00.690 \longrightarrow 00:05:03.480$ where we first consider the setting
- 114 00:05:03.480 --> 00:05:05.940 of an egocentric network randomized trial
- $115\ 00:05:05.940 \longrightarrow 00:05:08.190$ where at most one person per network
- $116\ 00:05:08.190 \longrightarrow 00:05:10.200$ can receive the intervention.
- $117\ 00:05:10.200 --> 00:05:11.610$ Then we extend our methods
- $118\ 00:05:11.610 \longrightarrow 00:05:13.410$ to consider cluster randomized trials
- $119\ 00:05:13.410 \longrightarrow 00:05:15.240$ where multiple participants per cluster
- $120\ 00:05:15.240 \longrightarrow 00:05:17.070$ can receive the intervention.
- $121\ 00:05:17.070 \longrightarrow 00:05:19.260$ And we also consider general settings
- $122\ 00:05:19.260 \longrightarrow 00:05:21.600$ where interference sets can be mismeasured
- $123\ 00{:}05{:}21.600 \dashrightarrow 00{:}05{:}24.993$ and the exposure is not necessarily randomized.
- 124 00:05:27.030 --> 00:05:28.800 So I'll begin with the first paper
- 125 00:05:28.800 --> 00:05:31.710 on egocentric network randomized trials.
- $126\ 00{:}05{:}31.710 \dashrightarrow 00{:}05{:}34.380$ So under this design, we have index participants
- $127\ 00:05:34.380 \longrightarrow 00:05:36.480$ who are recruited into this study
- $128\ 00:05:36.480 \longrightarrow 00:05:40.080$ and they're each asked to nominate a set of network members,
- $129\ 00:05:40.080 \longrightarrow 00:05:43.710$ which can be their drug injection partners or sex partners,
- $130\ 00:05:43.710 \longrightarrow 00:05:46.110$ and they form egocentric networks.
- $131\ 00{:}05{:}46.110 \longrightarrow 00{:}05{:}49.410$ And the index participants are the ones in the study
- $132\ 00:05:49.410 --> 00:05:52.110$ who are randomized to receive their intervention.
- $133\ 00:05:52.110 --> 00:05:56.550$ And examples of this are typically
- $134\ 00:05:56.550 \longrightarrow 00:05:59.310$ be peer education or behavioral-based.

- $135\ 00{:}05{:}59.310 \dashrightarrow 00{:}06{:}02.940$ And the index participants are asked to encourage
- $136\ 00:06:02.940 \longrightarrow 00:06:05.103$ behavioral change to their network members.
- $137\ 00:06:06.930 \longrightarrow 00:06:08.220$ So for some notation,
- 138 00:06:08.220 --> 00:06:10.860 we have participant ik being the i participant
- $139\ 00:06:10.860 \longrightarrow 00:06:12.810$ in the k network.
- $140\ 00{:}06{:}12.810 {\: -->\:} 00{:}06{:}16.350$ And we'll let i equal one denote the index participant
- $141\ 00:06:16.350 \longrightarrow 00:06:17.970$ in each network and I incur then one
- $142\ 00:06:17.970 --> 00:06:20.220$ denote the network members.
- $143\ 00:06:20.220 \longrightarrow 00:06:24.360$ We'll also define a network neighborhood for participant ik,
- 144 00:06:24.360 --> 00:06:26.010 which comprises of participants
- $145\ 00:06:26.010 \longrightarrow 00:06:28.293$ who share a network link with ik.
- $146\ 00{:}06{:}30.060 {\:{\mbox{--}}\!>}\ 00{:}06{:}33.180$ And then we also have a true membership matrix
- $147\ 00{:}06{:}33.180 \dashrightarrow 00{:}06{:}36.576$ which essentially represents whether a participant
- $148\ 00:06:36.576 \longrightarrow 00:06:39.360$ is a network member of a certain index.
- $149\ 00:06:39.360 \longrightarrow 00:06:43.443$ And we also have an intervention assignment indicator,
- $150\ 00:06:44.520 \longrightarrow 00:06:46.290$ which again the intervention is randomized
- $151\ 00:06:46.290 \longrightarrow 00:06:48.213$ and only received by the index member.
- 152 00:06:51.240 --> 00:06:54.480 So we'll let, throughout this dissertation,
- $153\ 00:06:54.480 --> 00:06:57.450$ represent an individual exposure to the intervention.
- $154\ 00{:}06{:}57.450 \dashrightarrow 00{:}07{:}01.470$ So because in here in NNRT, only index participants
- $155\ 00:07:01.470 --> 00:07:05.070$ who are randomized to treatment can receive the treatment
- $156\ 00{:}07{:}05.070$ --> $00{:}07{:}08.880$ and therefore, A is only equal to one for a treated index
- $157\ 00:07:08.880 \longrightarrow 00:07:11.913$ and A is equal to zero for everyone else.

 $158\ 00:07:12.810 \longrightarrow 00:07:15.600$ And so to define potential outcomes under interference,

 $159\ 00{:}07{:}15.600 \dashrightarrow 00{:}07{:}18.450$ we need to make assumptions on the interference structure.

 $160\ 00:07:18.450 \longrightarrow 00:07:22.260$ So here, we assume neighborhood interference

 $161\ 00:07:22.260 \longrightarrow 00:07:26.100$ with an exposure mapping function, which essentially says

162 00:07:26.100 --> 00:07:28.380 that I case potential outcome is determined

 $163\ 00:07:28.380 \longrightarrow 00:07:30.070$ by their own individual exposure

 $164~00{:}07{:}31.127 \dashrightarrow 00{:}07{:}33.870$ and the exposures of those in I case network neighborhood

165 00:07:33.870 --> 00:07:35.640 and not anyone outside of it,

 $166\ 00:07:35.640 \longrightarrow 00:07:38.230$ including participants from other networks

167 00:07:39.180 --> 00:07:42.030 and out of study individuals.

 $168\ 00{:}07{:}42.030$ --> $00{:}07{:}45.690$ So we further apply an exposure mapping function

 $169\ 00:07:45.690 \longrightarrow 00:07:47.880$ to this network neighborhood

 $170\ 00{:}07{:}47.880 {\: -->\:} 00{:}07{:}51.090$ and in this paper, we consider an exposure mapping function

 $171\ 00:07:51.090 --> 00:07:54.090$ defined by the number of treated neighbors.

172 00:07:54.090 --> 00:07:57.630 So under this assumption,

173 00:07:57.630 --> 00:08:01.170 ik's potential outcome is given by y indexed by A and G,

 $174\ 00:08:01.170 \longrightarrow 00:08:02.457$ which is their individual exposure

 $175\ 00:08:02.457 --> 00:08:06.000$ and the spillover exposure given by the number

 $176\ 00{:}08{:}06.000 \dashrightarrow 00{:}08{:}08.703$ of treated neighbors in their neighbor neighborhood.

 $177\ 00:08:11.010 \longrightarrow 00:08:14.340$ So this figure is a representation

 $178\ 00:08:14.340 \longrightarrow 00:08:16.290$ of two networks where in order

179 00:08:16.290 --> 00:08:19.050 to define the spillover exposure,

 $180\ 00:08:19.050 --> 00:08:21.060$ we further make the assumption

 $181\ 00:08:21.060 --> 00:08:23.460$ that the networks are not overlapping.

182 00:08:23.460 --> 00:08:25.470 And so this means that index participants

- $183\ 00:08:25.470 \longrightarrow 00:08:27.720$ cannot be connected amongst themselves
- $184\ 00:08:27.720 --> 00:08:29.610$ and network members can only be connected
- $185\ 00:08:29.610 \longrightarrow 00:08:31.503$ to one index participant.
- $186\ 00:08:32.490 \dashrightarrow 00:08:35.550$ And by making this assumption we can obtain G
- $187~00{:}08{:}35.550 \dashrightarrow 00{:}08{:}38.790$ by multiplying M and R, which is the membership matrix
- $188\ 00:08:38.790 \longrightarrow 00:08:40.560$ and intervention assignment.
- $189\ 00:08:40.560 \longrightarrow 00:08:42.000$ And so the spillover exposure
- $190\ 00:08:42.000 \longrightarrow 00:08:43.770$ for each participant is only determined
- $191\ 00:08:43.770 \longrightarrow 00:08:47.160$ by whether they are connected to a treated index member,
- $192\ 00:08:47.160 \longrightarrow 00:08:49.023$ which is shown in this figure.
- $193\ 00:08:52.890 \longrightarrow 00:08:55.590$ So the causal estimate of interest in this paper,
- $194\ 00:08:55.590 \longrightarrow 00:08:58.320$ the average spillover effect which is the impact
- $195\ 00:08:58.320 \longrightarrow 00:09:01.050$ of the intervention on the network members.
- $196\ 00:09:01.050 \longrightarrow 00:09:03.180$ And we here, we define the spillover effect
- 197 00:09:03.180 --> 00:09:05.880 as a risk difference and as a risk ratio.
- 198 00:09:05.880 --> 00:09:08.190 And under assumptions of positivity
- $199\ 00:09:08.190 \dashrightarrow 00:09:12.630$ and unconfoundedness which is guaranteed in the ENRG design
- 200 00:09:12.630 --> 00:09:14.733 under perfect treatment compliance,
- $201\ 00:09:20.190 \longrightarrow 00:09:21.750$ we can identify the spillover effects
- $202\ 00:09:21.750 \longrightarrow 00:09:24.780$ using observed outcomes and estimate them
- $203\ 00:09:24.780 \longrightarrow 00:09:26.380$ using observed outcomes as well.
- $204\ 00:09:29.400 \longrightarrow 00:09:32.670$ So under the unconfoundedness assumption,
- $205\ 00:09:32.670 \dashrightarrow > 00:09:36.390$ if we had data on the true exposures, we would estimate
- $206\ 00{:}09{:}36.390 \dashrightarrow 00{:}09{:}40.097$ the spillover effect using sample average estimators
- $207\ 00:09:40.097 \longrightarrow 00:09:42.333$ using the two by two table at the top.
- $208\ 00:09:43.230$ --> 00:09:45.690 The issue with this is that in ENRTs, the networks

- $209\ 00:09:45.690 \longrightarrow 00:09:47.490$ that are observed, which are the ones
- 210 00:09:47.490 --> 00:09:49.350 that are collected at study baseline,
- $211\ 00:09:49.350 \dashrightarrow 00:09:51.690$ they may not represent the true connections that take place
- $212\ 00:09:51.690 \longrightarrow 00:09:53.580$ during the study period.
- $213\ 00:09:53.580 --> 00:09:56.340$ So for example, a network member can fall out of touch
- $214\ 00{:}09{:}56.340 \dashrightarrow 00{:}09{:}58.710$ with the index participant that they enrolled with
- $215\ 00{:}09{:}58.710 \dashrightarrow 00{:}10{:}03.540$ and they can also be friend another index of another network.
- $216\ 00:10:03.540 \longrightarrow 00:10:06.150$ And so under these observed networks,
- $217\ 00:10:06.150 --> 00:10:10.020$ there's spillover exposure may also be misclassified.
- $218\ 00:10:10.020$ --> 00:10:13.560 And using these misclassified spillover exposures,
- $219\ 00:10:13.560 \longrightarrow 00:10:15.930$ we will instead estimate the spillover effect
- $220\ 00{:}10{:}15{.}930 \dashrightarrow 00{:}10{:}20{.}220$ using the two by two table at the bottom of this slide.
- 221 00:10:20.220 --> 00:10:22.620 And we show that the estimated spillover effects
- $222\ 00:10:22.620 \longrightarrow 00:10:24.873$ under this table would be biased.
- $223\ 00{:}10{:}27.030 \dashrightarrow 00{:}10{:}29.430$ And here's a representation of the types
- $224\ 00:10:29.430 \dashrightarrow 00:10:33.720$ of network misclassification that can occur in ENRT.
- $225\ 00:10:33.720 \longrightarrow 00:10:35.517$ So here, the black links represent
- $226\ 00:10:35.517 \longrightarrow 00:10:38.640$ the correctly measured network ties.
- $227\ 00{:}10{:}38.640 \dashrightarrow 00{:}10{:}42.420$ The blue links represent network links that are observed
- 228 00:10:42.420 --> 00:10:44.130 but are in fact not true.
- $229\ 00{:}10{:}44.130 \dashrightarrow 00{:}10{:}47.130$ And the red links represent the ones that are not observed
- $230\ 00:10:47.130 \longrightarrow 00:10:50.160$ but in fact occur during the study period.
- $231\ 00:10:50.160 \longrightarrow 00:10:54.600$ And because of these types of misclassification

- 232 00:10:54.600 --> 00:10:56.490 person's truth, spillover exposure
- $233\ 00:10:56.490 \longrightarrow 00:10:58.540$ can be different from their observed one.
- $234\ 00:11:02.264 \longrightarrow 00:11:04.410$ In this paper, we further assume
- 235 00:11:04.410 --> 00:11:06.690 non-differential misclassification,
- $236\ 00:11:06.690 \longrightarrow 00:11:08.640$ which is that the misclassification process
- 237 00:11:08.640 --> 00:11:10.683 doesn't depend on potential outcomes.
- $238\ 00{:}11{:}11.550 \dashrightarrow 00{:}11{:}15.030$ So under this assumption we derive an expression
- 239 00:11:15.030 --> 00:11:17.460 for the bias using four parameters,
- 240 00:11:17.460 --> 00:11:20.100 which are the baseline Malcolm rate,
- $241\ 00{:}11{:}20.100 \dashrightarrow 00{:}11{:}24.510$ the true spillover risk ratio, PM, which is the probability
- $242\ 00:11:24.510 \longrightarrow 00:11:27.870$ of being classified into the correct network as well as PR,
- $243\ 00{:}11{:}27.870 \dashrightarrow 00{:}11{:}30.513$ which is the intervention allocation probability.
- 244 00:11:31.500 --> 00:11:33.930 And using these expressions we can show
- 245 00:11:33.930 --> 00:11:36.240 that there's no bias when PM is one,
- $246\ 00:11:36.240 \longrightarrow 00:11:38.460$ which is when everyone is correctly classified
- $247\ 00:11:38.460 \longrightarrow 00:11:40.173$ due to the correct network.
- $248\ 00{:}11{:}41.220$ --> $00{:}11{:}44.220$ We can also show that the bias is always towards the null
- $249\ 00{:}11{:}44.220 \dashrightarrow 00{:}11{:}47.010$ under the non differential misclassification assumption.
- $250\ 00:11:47.010 --> 00:11:49.680$ So the ASP would always be underestimated
- $251\ 00:11:49.680 \longrightarrow 00:11:50.760$ under this assumption
- $252\ 00:11:50.760 \longrightarrow 00:11:53.193$ if spillover exposures were misclassified.
- 253 00:11:56.610 --> 00:11:58.590 So in order to correct for this bias,
- $254\ 00:11:58.590 \longrightarrow 00:12:01.380$ we use a validation study.
- $255\ 00:12:01.380 \longrightarrow 00:12:03.210$ So again, this is where the true network
- $256\ 00:12:03.210 \longrightarrow 00:12:05.100$ or spillover exposure is measured
- $257\ 00:12:05.100 --> 00:12:06.690$ alongside the mismeasured ones
- $258\ 00:12:06.690 \longrightarrow 00:12:09.030$ for a subsample of the main study.

- $259\ 00:12:09.030 \longrightarrow 00:12:12.180$ And then in this paper, we estimate the sensitivity
- $260\ 00{:}12{:}12.180 \dashrightarrow 00{:}12{:}14.880$ and specificity of spillover exposure classification
- $261\ 00{:}12{:}14.880 \dashrightarrow 00{:}12{:}18.090$ among network members and we assume that the parameters
- $262\ 00:12:18.090 \longrightarrow 00:12:20.040$ that are estimated in the validation study
- 263 00:12:20.040 --> 00:12:22.353 is generalizable to the main study.
- $264\ 00:12:24.180 --> 00:12:25.920$ We can show that the sensitivity
- $265~00{:}12{:}25.920 \dashrightarrow 00{:}12{:}30.800$ and specificity can be expressed as functions of PM and PR
- $266\ 00:12:31.830 \longrightarrow 00:12:35.370$ where the intuition is that if a participant
- $267\ 00{:}12{:}35.370 \dashrightarrow 00{:}12{:}38.040$ or if a network member is observed to be connected
- $268~00{:}12{:}38.040 \dashrightarrow 00{:}12{:}42.600$ to a treated index, given that there really are connected
- $269\ 00:12:42.600 \longrightarrow 00:12:43.947$ to a treated index,
- $270\ 00:12:43.947 \longrightarrow 00:12:47.190$ this can be because they are correctly classified
- 271 00:12:47.190 --> 00:12:49.110 or it could be because they were misclassified
- 272 00:12:49.110 --> 00:12:50.130 but still connected
- $273\ 00:12:50.130 --> 00:12:52.473$ to a treated index just from another network.
- $274\ 00:12:53.610 \longrightarrow 00:12:55.290$ We can also estimate the sensitivity
- $275\ 00:12:55.290 \dashrightarrow 00:12:58.260$ and specificity using the two by two table.
- $276\ 00{:}12{:}58.260 \dashrightarrow 00{:}13{:}01.680$ Here, when we assume that the misclassification process
- $277\ 00:13:01.680 --> 00:13:03.657$ doesn't depend on covariate.
- $278\ 00:13:07.200 \longrightarrow 00:13:10.470$ So we propose three estimators in this paper.
- $279\ 00{:}13{:}10.470 \dashrightarrow 00{:}13{:}13.200$ The first is called the matrix method estimator.
- $280\ 00:13:13.200 \longrightarrow 00:13:17.460$ And here, this estimator takes the estimated sensitivity
- 281 00:13:17.460 --> 00:13:19.890 and specificity from the validation study
- 282 00:13:19.890 --> 00:13:22.380 to bias correct accounts observed
- $283\ 00:13:22.380 \longrightarrow 00:13:24.720$ from the two by two table in the main study.

- $284\ 00{:}13{:}24.720 \dashrightarrow 00{:}13{:}27.480$ And this would be the form of the bias corrected estimators
- $285\ 00:13:27.480 \longrightarrow 00:13:29.310$ for this spillover effect.
- $286\ 00:13:29.310 \longrightarrow 00:13:30.990$ And we can obtain its variance
- $287\ 00:13:30.990 \longrightarrow 00:13:33.330$ by the multivariate delta method.
- $288\ 00{:}13{:}33.330 {\:{\mbox{--}}\!>} 00{:}13{:}36.120$ And if we believe that there is clustering in the study,
- $289\ 00{:}13{:}36.120 \dashrightarrow 00{:}13{:}40.290$ we can also adjust for this by a design effect inflation
- 290 00:13:40.290 --> 00:13:42.630 or we can perform network bootstrapping
- $291\ 00:13:42.630 --> 00:13:44.480$ where we re-sample networks as whole.
- 292 00:13:46.110 --> 00:13:48.960 And we know that to use this method
- $293\ 00:13:48.960 \longrightarrow 00:13:51.750$ and for this method to perform well,
- $294\ 00{:}13{:}51.750 \dashrightarrow 00{:}13{:}54.510$ there needs to be constraints on the value of sensitivity
- $295\ 00:13:54.510 \longrightarrow 00:13:57.510$ and specificity for the estimator to be stable
- $296\ 00:13:57.510 \longrightarrow 00:14:00.603$ and to avoid estimating negative cell counts.
- 297 00:14:02.130 --> 00:14:05.220 So when these constraints are not met,
- $298\ 00{:}14{:}05.220$ --> $00{:}14{:}08.650$ we can instead consider an inverse matrix method estimator
- 299 00:14:09.720 --> 00:14:13.320 which corrects the cell counts in the two at two table
- $300\ 00:14:13.320 \longrightarrow 00:14:14.940$ in the main study using the positive
- 301 00:14:14.940 --> 00:14:16.320 and negative predictive values
- 302 00:14:16.320 --> 00:14:19.020 instead of the sensitivity and specificity.
- $303\ 00:14:19.020 --> 00:14:22.380$ And this method uses the PPV and MPV
- $304\ 00{:}14{:}22.380 \dashrightarrow 00{:}14{:}25.860$ estimated separately for those with and without the outcome.
- $305~00{:}14{:}25.860 \dashrightarrow 00{:}14{:}28.320$ And therefore the matrix method estimated
- $306\ 00:14:28.320 \longrightarrow 00:14:31.230$ may be more efficient relative to this estimator
- $307\ 00:14:31.230 \longrightarrow 00:14:32.190$ if the outcome is rare
- $308\ 00:14:32.190 --> 00:14:33.850$ and the validation study is small
- $309\ 00:14:36.570 --> 00:14:38.940$ and the last estimator we considered

- $310\ 00:14:38.940 --> 00:14:41.880$ was a likelihood-based estimator
- $311\ 00:14:41.880 \longrightarrow 00:14:43.290$ because while the matrix
- $312\ 00{:}14{:}43.290 \dashrightarrow 00{:}14{:}46.680$ and inverse matrix estimators are easily implemented,
- $313\ 00{:}14{:}46.680 --> 00{:}14{:}49.470$ there's no clear way of directly incorporating the effect
- $314\ 00:14:49.470 \longrightarrow 00:14:51.570$ of clustering into the entrance.
- $315\ 00{:}14{:}51.570 \dashrightarrow 00{:}14{:}55.140$ And therefore, we can specify an outcome model
- $316\ 00:14:55.140 --> 00:14:57.100$ including a network random effect
- 317 00:14:59.734 --> 00:15:02.100 to account for clustering by networks
- 318 00:15:02.100 --> 00:15:06.000 and using the likelihood specified here,
- $319\ 00:15:06.000 --> 00:15:08.593$ we can obtain the MLE of the ASP
- $320\ 00:15:09.930 \longrightarrow 00:15:12.930$ and its variance by the inverse
- 321 00:15:12.930 --> 00:15:14.830 of the observation information matrix.
- 322 00:15:18.870 --> 00:15:22.936 Right, so I'll go over our application
- 323 00:15:22.936 --> 00:15:25.770 of our methods using the HPTN 037 study,
- $324\ 00{:}15{:}25.770 \dashrightarrow 00{:}15{:}29.640$ which was an ENRT that was conducted in Philadelphia
- 325 00:15:29.640 --> 00:15:31.380 and Chiang Mai Thailand
- $326\ 00:15:31.380 \longrightarrow 00:15:34.170$ where in the study indexes were randomized
- $327\ 00{:}15{:}34.170 \dashrightarrow 00{:}15{:}36.240$ to receive an intervention that consisted
- $328\ 00{:}15{:}36.240 {\: -->\:} 00{:}15{:}39.310$ of a peer education training where they were encouraged
- 329 00:15:41.670 --> 00:15:43.950 to disseminate HIV knowledge
- $330\ 00:15:43.950 \longrightarrow 00:15:46.740$ and injection sexual risk reduction behaviors
- 331 00:15:46.740 --> 00:15:48.183 with their network members.
- $332\ 00{:}15{:}49.160 \dashrightarrow 00{:}15{:}53.130$ In the study, we were interested in looking at the effect
- $333\ 00{:}15{:}53.130 \dashrightarrow 00{:}15{:}57.020$ of the intervention on any self-reported HIV risk behaviors
- $334\ 00:15:57.020 \longrightarrow 00:15:59.643$ at one year after study enrollment.
- $335\ 00:16:00.990 \longrightarrow 00:16:02.820$ Here, we define G star,

- $336\ 00:16:02.820 \longrightarrow 00:16:05.010$ which is they observed spillover exposure
- $337\ 00{:}16{:}05.010 \dashrightarrow 00{:}16{:}07.110$ based on the intervention assigned to the networks
- $338\ 00:16:07.110 \longrightarrow 00:16:09.480$ and receive by their index members.
- $339\ 00:16:09.480 --> 00:16:12.870$ And we define G the true spillover exposure
- $340\ 00:16:16.500 --> 00:16:18.810$ based on an exposure contamination survey
- 341 00:16:18.810 --> 00:16:21.990 that was taken at six months post baseline.
- $342\ 00:16:21.990 --> 00:16:24.120$ So in the survey, participants were asked
- $343\ 00:16:24.120 \longrightarrow 00:16:28.230$ to recall five specific terminologies associated
- $344\ 00:16:28.230 \longrightarrow 00:16:30.780$ with the intervention training.
- $345\ 00{:}16{:}30.780 \dashrightarrow 00{:}16{:}33.150$ So we suppose that if a network member were able
- $346\ 00:16:33.150 \longrightarrow 00:16:36.090$ to recall any of these five terms, that they were exposed
- $347\ 00:16:36.090 \longrightarrow 00:16:38.430$ to the intervention through a treated index
- $348\ 00:16:38.430 \longrightarrow 00:16:39.690$ and if they weren't able to recall
- $349\ 00:16:39.690 \longrightarrow 00:16:42.690$ any of the terms, then they weren't exposed.
- 350 00:16:42.690 --> 00:16:44.940 And because there was a possibility
- $351\ 00:16:44.940 \longrightarrow 00:16:48.300$ that network members may be exposed to the training
- $352\ 00:16:48.300 \longrightarrow 00:16:50.580$ but just didn't remember any of the terms
- $353\ 00:16:50.580 --> 00:16:53.340$ or that there may be network members
- $354\ 00:16:53.340 \longrightarrow 00:16:55.590$ who were not exposed to the training
- $355\ 00:16:55.590 --> 00:16:57.360$ but just said that they remember terms
- 356 00:16:57.360 --> 00:16:59.460 because of social desirability,
- $357\ 00:16:59.460 --> 00:17:01.170$ we only included network members
- 358~00:17:01.170 --> 00:17:04.140 who recalled the positive control term which was exposed
- $359\ 00:17:04.140 --> 00:17:07.500$ to every body regardless of their randomization arm
- $360\ 00:17:07.500 \longrightarrow 00:17:09.480$ and none of the negative control terms,
- 361~00:17:09.480 --> 00:17:13.830 which none of the participants were supposed to know

- $362\ 00:17:13.830 \longrightarrow 00:17:14.880$ that only these participants
- $363\ 00:17:14.880 \longrightarrow 00:17:16.320$ were included in the validation study
- $364\ 00:17:16.320 --> 00:17:18.540$ so we can more accurately estimate
- $365\ 00:17:18.540 \longrightarrow 00:17:20.313$ the sensitivity and specificity.
- $366\ 00:17:22.860 \longrightarrow 00:17:27.000$ So here are the effects of the intervention
- $367\ 00:17:27.000 \longrightarrow 00:17:28.950$ or the spillover effects of the intervention
- $368\ 00{:}17{:}28.950 \dashrightarrow 00{:}17{:}33.950$ on risk behaviors were from the validation study we see
- $369\ 00:17:34.050 \longrightarrow 00:17:36.420$ that there was indeed some degree
- $370~00{:}17{:}36.420 \dashrightarrow 00{:}17{:}40.620$ of network misclassification where the sensitivity was 60%
- $371\ 00:17:40.620 \longrightarrow 00:17:43.800$ and specificity was 79%.
- 372 00:17:43.800 --> 00:17:47.100 So then intent-to-treat estimator uses G star,
- $373\ 00:17:47.100 --> 00:17:49.950$ which is the intervention assigned to the networks.
- $374\ 00:17:49.950 \longrightarrow 00:17:51.270$ And we do see already
- $375\ 00:17:51.270 \longrightarrow 00:17:53.940$ that there was significant spillover effect
- $376\ 00:17:53.940 \longrightarrow 00:17:57.030$ of the intervention on reducing risk behaviors.
- 377 00:17:57.030 --> 00:17:59.010 And this effect was amplified
- $378\ 00:17:59.010 \longrightarrow 00:18:01.653$ after applying our bias correction method.
- $379~00{:}18{:}03.480 \dashrightarrow 00{:}18{:}07.170$ So we applied the matrix method estimator as well as the MLE
- $380\ 00:18:07.170 \longrightarrow 00:18:09.760$ and the inverse matrix method
- $381\ 00:18:10.939 \longrightarrow 00:18:13.590$ was not an ideal choice in this study
- 382 00:18:13.590 --> 00:18:15.420 because of our small validation study
- $383\ 00:18:15.420 --> 00:18:17.730$ and the number of participants
- 384 00:18:17.730 --> 00:18:19.580 who had the outcome within the study.
- $385\ 00:18:20.730 \longrightarrow 00:18:25.590$ Here, we also compared several standard errors were first,
- $386~00:18:25.590 \longrightarrow 00:18:27.030$ we consider standard standards obtained
- $387\ 00:18:27.030 --> 00:18:30.090$ from the delta method and those inflated
- $388\ 00:18:30.090 \longrightarrow 00:18:31.770$ by the design effect.

- $389\ 00:18:31.770 \longrightarrow 00:18:33.990$ And we see that the confidence intervals here
- $390\ 00:18:33.990 \longrightarrow 00:18:36.480$ were pretty wide, which is due
- $391\ 00:18:36.480 \longrightarrow 00:18:39.510$ to the small validation study sample size.
- $392\ 00:18:39.510 \longrightarrow 00:18:41.910$ However, when we consider network bootstrapping
- $393\ 00:18:41.910 \longrightarrow 00:18:43.770$ or the likelihood base method,
- $394\ 00:18:43.770 --> 00:18:46.770$ we see that the confidence interval significantly narrowed
- 395~00:18:46.770 --> 00:18:49.983 and we were able to see a significant spillover effect.
- 396 00:18:53.820 --> 00:18:55.800 So as a summary of this first paper,
- $397\ 00:18:55.800 \longrightarrow 00:18:58.470$ we proposed several bias correction estimators
- $398\ 00:18:58.470 \longrightarrow 00:19:00.480$ for this spillover effect
- $399\ 00:19:00.480$ --> 00:19:04.290 to address network misclassification and NRTs.
- $400\ 00{:}19{:}04.290 {\:{\circ}{\circ}{\circ}}>00{:}19{:}07.050$ So our methods here assume that both the exposure
- $401\ 00:19:07.050 \longrightarrow 00:19:08.730$ and outcome are binary measures
- $402\ 00:19:08.730 --> 00:19:12.480$ and we did not consider covariate adjustment
- $403\ 00:19:12.480 \longrightarrow 00:19:15.060$ because the intervention is randomized.
- $404\ 00:19:15.060 \longrightarrow 00:19:17.610$ And so as a segue to the second paper,
- $405\ 00:19:17.610 \longrightarrow 00:19:19.620$ we will be developing methods
- 406 00:19:19.620 --> 00:19:22.710 for non-binary exposures outcomes
- $407\ 00:19:22.710 \longrightarrow 00:19:25.620$ as well as allowing for covariate adjustment.
- $408\ 00:19:25.620 --> 00:19:27.390$ And we develop these methods in the setting
- $409\ 00:19:27.390 \longrightarrow 00:19:29.920$ of cluster randomized trials.
- $410\ 00:19:34.766 \longrightarrow 00:19:37.910$ So causal inference in cluster randomized trials
- 411 00:19:37.910 --> 00:19:41.100 in CRTs often rely on the assumption
- 412 00:19:41.100 --> 00:19:42.510 of partial interference,
- $413\ 00:19:42.510 \longrightarrow 00:19:45.220$ which is that participants are separated
- 414 00:19:46.115 --> 00:19:47.460 into non-overlapping clusters
- $415\ 00:19:47.460 \longrightarrow 00:19:49.890$ and interference is assumed

- $416\ 00:19:49.890 \longrightarrow 00:19:52.050$ to be only contained within these clusters
- $417\ 00:19:52.050 \longrightarrow 00:19:53.910$ and not across clusters.
- 418 00:19:53.910 --> 00:19:56.100 And this assumption is typically made
- 419 00:19:56.100 -> 00:20:00.543 because there is an absence of social network data and CRTs.
- $420\ 00:20:02.220$ --> 00:20:04.510 So interference sets define other departure
- $421\ 00:20:05.613 \longrightarrow 00:20:06.870$ interference assumptions are usually given
- $422\ 00:20:06.870 \longrightarrow 00:20:09.780$ by the randomization clusters in the trial.
- $423\ 00:20:09.780 \longrightarrow 00:20:12.250$ So this can be villages or communities
- $424\ 00:20:15.566 --> 00:20:17.130$ and the interference says they're given
- $425\,00:20:17.130\,\text{--}{>}\,00:20:20.280$ by the randomization clusters can be measured with there
- $426\ 00{:}20{:}20{:}20{:}20{:}23{:}970$ because they might be a lot larger than the true networks
- $427\ 00:20:23.970 \longrightarrow 00:20:26.370$ if they were considered to be whole communities.
- $428\ 00:20:26.370 \longrightarrow 00:20:28.020$ And also interactions can exist
- $429\ 00:20:28.020 \longrightarrow 00:20:29.883$ across these communities as well.
- $430\ 00{:}20{:}30.720 \dashrightarrow 00{:}20{:}34.590$ So this figure was taken from a file genetic analysis
- $431\ 00{:}20{:}34.590 \dashrightarrow 00{:}20{:}39.590$ from BCPP where BCPP was in HIV prevention CRT
- 432 $00:20:39.900 \longrightarrow 00:20:43.620$ that was conducted in 30 communities in Botswana.
- $433\ 00:20:43.620 \longrightarrow 00:20:47.640$ And so the randomization clusters were communities
- $434\ 00{:}20{:}47.640 \dashrightarrow 00{:}20{:}51.430$ but from the phylogenetic analysis where they sequenced
- $435\ 00:20:52.380 \longrightarrow 00:20:56.590\ HIV\ genes\ viral\ sequences$
- $436\ 00:20:57.840 \longrightarrow 00:21:00.330$ that they saw that the viral transmission chains obtained
- $437\ 00:21:00.330 \longrightarrow 00:21:03.720$ from the sequences, the majority of them
- 438 00:21:03.720 --> 00:21:06.300 actually crossed two or more communities,
- $439\ 00:21:06.300 \longrightarrow 00:21:07.320$ which was an indication

- 440 00:21:07.320 --> 00:21:09.840 of high-end cluster mixing in this study.
- 441 00:21:09.840 --> 00:21:11.490 And interference says that are defined
- $442\ 00{:}21{:}11.490 --> 00{:}21{:}14.883$ just by communities would be misspecified in this case.
- 443 00:21:18.180 --> 00:21:20.387 So here we again have participant ik
- $444\ 00:21:20.387 --> 00:21:24.300$ as the i participant indicate cluster.
- $445\ 00{:}21{:}24.300 \dashrightarrow 00{:}21{:}28.260$ We have script one and to denote the study sample.
- $446\ 00{:}21{:}28.260$ --> $00{:}21{:}33.260$ In this study, we first consider a two-stage CRT
- 447 00:21:33.330 --> 00:21:35.640 where clusters are first randomized
- $448\ 00:21:35.640 \longrightarrow 00:21:38.430$ to an intervention allocation strategy.
- $449\ 00{:}21{:}38.430 \dashrightarrow 00{:}21{:}41.580$ Here, we consider strategies alpha one and alpha two
- $450\ 00:21:41.580 \longrightarrow 00:21:44.040$ and alpha one and alpha two are probabilities.
- 451 00:21:44.040 --> 00:21:46.020 And under a balanced design,
- $452\ 00{:}21{:}46.020 \dashrightarrow 00{:}21{:}48.450$ half of the clusters would be assigned to alpha one
- $453\ 00{:}21{:}48.450 \dashrightarrow 00{:}21{:}51.270$ and half would be assigned to alpha alpha two.
- 454 00:21:51.270 --> 00:21:55.140 Then after the first stage randomization,
- $455\ 00:21:55.140 --> 00:21:56.670$ participants within these clusters
- $456\ 00{:}21{:}56.670 \dashrightarrow 00{:}21{:}59.280$ would be randomized to receive the intervention
- $457\ 00:21:59.280 \longrightarrow 00:22:01.050$ with the probability equal to the one
- $458\ 00:22:01.050 \longrightarrow 00:22:02.800$ that was assigned to their cluster.
- $459\ 00{:}22{:}03.900 \dashrightarrow 00{:}22{:}07.230$ And we'll extend our methods to consider general CRTs,
- 460 00:22:07.230 --> 00:22:09.840 which can be considered as a special case
- 461 00:22:09.840 --> 00:22:12.190 of a two-stage design where alpha one
- $462\ 00:22:12.190 \longrightarrow 00:22:14.880$ and alpha two are one and zero, which means
- $463\ 00{:}22{:}14.880 \dashrightarrow 00{:}22{:}18.060$ that clusters are randomized to intervention or control
- $464\ 00:22:18.060 \longrightarrow 00:22:20.400$ and there isn't a second stage randomization

- $465\ 00:22:20.400 \longrightarrow 00:22:21.813$ at the participant level.
- $466\ 00{:}22{:}25.110 \dashrightarrow 00{:}22{:}29.400$ So we again have to denote the individual exposure.
- $467\ 00:22:29.400 \longrightarrow 00:22:31.950$ And to define potential outcomes in the setting,
- $468\ 00:22:31.950 --> 00:22:35.340$ we first define a subset of the study sample
- $469\ 00:22:35.340 --> 00:22:39.153$ for participant ik and we denote this by script I.
- $470\ 00{:}22{:}40.110$ --> $00{:}22{:}42.850$ So here, we make the partial interference assumption
- 471 00:22:45.497 --> 00:22:48.030 where ik's potential outcome is influenced
- $472\ 00:22:48.030 \longrightarrow 00:22:51.690$ by their own exposure as well as the exposures
- $473\ 00:22:51.690 \longrightarrow 00:22:54.600$ of the participants within this subset
- $474\ 00:22:54.600 \longrightarrow 00:22:57.060$ and not anyone outside of this subset.
- $475\ 00{:}22{:}57.060 \dashrightarrow 00{:}23{:}00.041$ So because only the exposures of the participants
- $476\ 00:23:00.041 --> 00:23:03.910$ will affect the outcome of ik, we call this subset
- $477\ 00:23:04.864 \longrightarrow 00:23:06.164$ for ik's interference set.
- $478\ 00{:}23{:}07.514 \dashrightarrow 00{:}23{:}10.710$ And we can further apply an exposure mapping function
- $479\ 00:23:10.710 \longrightarrow 00:23:13.000$ to this interference set
- $480\ 00{:}23{:}13.920 \dashrightarrow 00{:}23{:}16.980$ to obtain a scaler quantity of a spillover exposure.
- 481 00:23:16.980 --> 00:23:20.463 Here, we consider stratify interference,
- $482\ 00:23:21.600 --> 00:23:24.540$ which essentially assumes that spillover occurs
- $483\ 00:23:24.540 \longrightarrow 00:23:26.880$ through the proportion of treated participants
- $484\ 00{:}23{:}26.880 \dashrightarrow 00{:}23{:}30.390$ in the interference set regardless of who they are.
- $485\ 00{:}23{:}30.390 \dashrightarrow 00{:}23{:}33.540$ So the spillover exposure would be given by disproportion
- $486\ 00:23:33.540 \longrightarrow 00:23:34.650$ and as in the first paper,
- $487\ 00:23:34.650 --> 00:23:37.893$ we can index potential outcomes by A and G.
- $488\ 00:23:41.490 --> 00:23:44.700$ Here, we consider four causal effects
- 489 00:23:44.700 --> 00:23:47.730 which the individual effect, spillover effect,

- $490\ 00:23:47.730 \longrightarrow 00:23:49.443$ total effect and overall effect.
- $491\ 00:23:50.280 \longrightarrow 00:23:53.227$ The individual effect is the effect
- $492\ 00:23:53.227$ --> 00:23:57.240 of the individual exposure under a fixed spillover exposure.
- $493\ 00{:}23{:}57.240 --> 00{:}24{:}00.900$ And on the other hand, the spillover effect is the effect
- $494\ 00:24:00.900 \longrightarrow 00:24:05.100$ of the spillover exposure under a fixed individual exposure.
- 495 00:24:05.100 --> 00:24:07.320 And then the total effect is the effect
- $496~00{:}24{:}07.320 \dashrightarrow 00{:}24{:}10.170$ of having both an individual exposure to the intervention
- $497\ 00:24:10.170 --> 00:24:12.390$ and some degree of spillover exposure
- $498\ 00:24:12.390 \longrightarrow 00:24:15.060$ versus neither type of exposure.
- $499\ 00:24:15.060 --> 00:24:17.280$ And then the overall effect compares the effect
- 500 00:24:17.280 --> 00:24:19.390 of being assigned to a cluster randomized
- $501~00{:}24{:}20.777 \dashrightarrow 00{:}24{:}23.377$ to treatment allocation strategy alpha versus alpha.
- $502\ 00:24:27.150 \longrightarrow 00:24:31.650$ So these causal effects can again be identified
- $503\ 00{:}24{:}31.650 \dashrightarrow 00{:}24{:}34.663$ under the assumption the identifying assumptions
- $504\ 00:24:34.663 \longrightarrow 00:24:36.900$ we made in the paper earlier
- $505\ 00:24:36.900 \longrightarrow 00:24:38.130$ or as in the first paper,
- 506 00:24:38.130 --> 00:24:41.520 which were the unconfounded assumption
- 507 00:24:41.520 --> 00:24:44.430 which would hold under a two-stage design
- 508 00:24:44.430 --> 00:24:47.190 given perfect treatment compliance.
- $509\ 00:24:47.190 --> 00:24:48.600$ And we estimate these effects
- $510\ 00:24:48.600 \longrightarrow 00:24:52.350$ using a regression-based estimation approach,
- $511\ 00:24:52.350 \longrightarrow 00:24:53.183$ which is consistent
- $512\ 00:24:53.183 \longrightarrow 00:24:55.860$ and efficient under a correctly specified model
- $513\ 00:24:55.860 \longrightarrow 00:24:57.460$ for the potential outcome.
- $514\ 00{:}24{:}58.560 {\:{\mbox{--}}}{>}\ 00{:}25{:}03.560$ So here, we consider an outcome model in this form

- $515\ 00{:}25{:}05.962 \dashrightarrow 00{:}25{:}08.747$ where we include a cluster random effect to account
- $516\ 00:25:10.050 \longrightarrow 00:25:12.147$ for the effect of clustering and the inference
- 517 00:25:12.147 --> 00:25:14.940 and we also have an interaction between A and G
- $518\ 00:25:14.940 \longrightarrow 00:25:18.150$ so that we can allow the individual effect to vary
- 519 00:25:18.150 --> 00:25:21.753 with G and the spillover effect to vary with A.
- $520\ 00:25:23.070$ --> 00:25:27.270 So once we have the estimated coefficients from this model,
- $521\ 00:25:27.270 \longrightarrow 00:25:29.760$ we can estimate causal effects
- 522 00:25:29.760 --> 00:25:31.953 using these estimated coefficients.
- 523 00:25:35.820 --> 00:25:37.890 So again, in CRTs,
- $524~00{:}25{:}37.890 \dashrightarrow 00{:}25{:}40.800$ because we don't have data on social connections
- $525\ 00:25:40.800 \longrightarrow 00:25:43.200$ and when we consider interference to be given
- $526~00{:}25{:}43.200 \dashrightarrow 00{:}25{:}46.950$ by randomization clusters, they can be measured with error.
- $527\ 00:25:46.950 --> 00:25:48.750$ And as a consequence, this spillover exposure
- $528\ 00:25:48.750 \longrightarrow 00:25:51.450$ can also be measured with error.
- $529\ 00:25:51.450 \longrightarrow 00:25:53.490$ So we have shown in this paper
- 530~00:25:53.490 --> 00:25:57.330 that when the outcome model is fit with G star instead of G,
- $531\ 00:25:57.330 \longrightarrow 00:25:59.610$ the estimated model coefficient will be biased
- $532\ 00:25:59.610 \longrightarrow 00:26:01.440$ and the causal effects estimated
- $533\ 00:26:01.440 --> 00:26:04.503$ with these bias coefficients were therefore also be biased.
- $534\ 00:26:07.707 \longrightarrow 00:26:09.810$ And to correct for the bias
- 535 00:26:09.810 --> 00:26:11.786 in these regression coefficients,
- $536\ 00:26:11.786 \longrightarrow 00:26:15.210$ we apply a regression calibration approach
- $537\ 00:26:15.210 --> 00:26:17.310$ which is developed under the assumption
- $538\ 00:26:17.310 --> 00:26:19.920$ that the measurement error is additive

- $539~00{:}26{:}19.920 \dashrightarrow 00{:}26{:}22.020$ and also the non differential measurement error
- $540\ 00:26:22.020 \longrightarrow 00:26:23.433$ as in the previous paper.
- 541 00:26:24.810 --> 00:26:27.180 So to apply this method,
- $542\ 00:26:27.180 \longrightarrow 00:26:29.520$ we will first regress the outcome
- $543~00{:}26{:}29.520 \dashrightarrow 00{:}26{:}31.200$ on the mismeasured exposure
- $544\ 00:26:31.200 \longrightarrow 00:26:33.660$ in the main study as we would
- 545 00:26:33.660 --> 00:26:36.240 under the intent-to-treat analysis.
- $546\ 00:26:36.240 \longrightarrow 00:26:37.425$ In the validation study
- $547\ 00:26:37.425 \longrightarrow 00:26:40.830$ because we assumed that the measurement error is additive,
- $548\ 00{:}26{:}40.830 {\: -->\:} 00{:}26{:}45.300$ we fit a linear measurement error model of the true exposure
- $549\ 00:26:45.300 \longrightarrow 00:26:49.020$ given the mismeasured spillover exposure.
- 550~00:26:49.020 --> 00:26:51.180 And then we can obtain bias corrected regression
- $551\ 00:26:51.180 \longrightarrow 00:26:53.479$ coefficients using the coefficients
- $552\ 00:26:53.479 --> 00:26:56.070$ obtained from these two models.
- 553 00:26:56.070 --> 00:26:58.800 And we can obtain the variance
- $554\ 00:26:58.800 \longrightarrow 00:27:01.713$ of these corrected coefficients using the delta.
- $555\ 00:27:05.730 \longrightarrow 00:27:09.030$ We can also extend this approach to account
- $556\ 00:27:09.030 --> 00:27:10.680$ for covariate adjustment
- $557\ 00:27:10.680 --> 00:27:12.630$ and there may be several reasons why we need
- $558\ 00:27:12.630 \longrightarrow 00:27:14.550$ to adjust for covariates.
- 559~00:27:14.550 --> 00:27:18.390 First, if we step out of the two stage CRT setting
- $560~00{:}27{:}18.390 \dashrightarrow 00{:}27{:}21.990$ and we consider a general CRT where intervention is work,
- $561\ 00:27:21.990$ --> 00:27:24.780 clusters are assigned to either intervention or control.
- $562\ 00{:}27{:}24.780 \dashrightarrow 00{:}27{:}28.590$ And a lot of public health studies that the interventions
- 563 00:27:28.590 --> 00:27:31.170 that are given to these clusters may be prone

- $564\ 00:27:31.170 \longrightarrow 00:27:33.090$ to non-compliance.
- 565 00:27:33.090 --> 00:27:36.360 And intervention uptake will depend
- $566\ 00:27:36.360 \longrightarrow 00:27:38.220$ on individual characteristics
- $567\ 00:27:38.220 \longrightarrow 00:27:41.460$ that may need to be accounted for.
- $568~00{:}27{:}41.460 \dashrightarrow 00{:}27{:}46.460$ So in the when there are confounders between the outcome
- $569\ 00:27:46.470 \longrightarrow 00:27:50.430$ and the individual exposure, A or G,
- $570\ 00{:}27{:}50.430 \dashrightarrow 00{:}27{:}53.380$ we would need to assume conditional unconfoundedness
- $571\ 00:27:54.870 \longrightarrow 00:27:57.000$ of the individual exposure exposures.
- $572\ 00:27:57.000 \longrightarrow 00:28:01.650$ So when there are covariates or confounders between Y and A,
- $573\ 00:28:01.650 \longrightarrow 00:28:05.250$ we would adjust them in the outcome model.
- 574~00:28:05.250 --> 00:28:09.360 And if they were confounders between Y and G,
- $575\ 00:28:09.360 \longrightarrow 00:28:11.130$ we would adjust for them in the outcome model
- $576\ 00:28:11.130 --> 00:28:13.833$ as well as in the measurement error model.
- $577\ 00{:}28{:}15.268$ --> $00{:}28{:}20.268$ We might need to also make the non-differential
- 578~00:28:20.340 --> 00:28:23.553 measurement error assumption conditional on covariates.
- $579~00{:}28{:}24.870 \dashrightarrow 00{:}28{:}27.420$ And in this case, because these covariates are related
- $580\ 00:28:27.420 \longrightarrow 00:28:29.820$ to the measurement error as well as the outcome.
- 581~00:28:29.820 --> 00:28:33.960 They would need to be adjusted in both models as well.
- $582\ 00:28:33.960 --> 00:28:36.390$ And lastly, we may only be able
- $583\ 00{:}28{:}36{.}390 \dashrightarrow 00{:}28{:}39{.}647$ to generalize the measurement error parameters estimated
- $584\ 00:28:39.647 --> 00:28:42.510$ in the validation study to the main study
- 585~00:28:42.510 --> 00:28:43.890 conditional on covariates.
- $586\ 00:28:43.890 \longrightarrow 00:28:46.140$ And in this case, we would adjust

- $587\ 00:28:46.140 \longrightarrow 00:28:49.023$ for these covariates as well.
- 588~00:28:49.860 --> 00:28:52.710 But regardless of the types of covariates that are adjusted
- $589\ 00:28:52.710 \longrightarrow 00:28:56.310$ for the regression calibration estimators
- $590\ 00:28:56.310 --> 00:28:57.807$ and variance estimators
- 591~00:29:01.328 --> 00:29:04.020 for the coefficients that are of interest that are used
- $592~00:29:04.020 \dashrightarrow 00:29:07.650$ to estimate the causal effects, they would not be changed
- $593\ 00:29:07.650 \longrightarrow 00:29:10.323$ as in the case without the variates.
- $594\ 00:29:13.650 --> 00:29:17.613$ We've applied our methods to the BCPP study,
- 595 00:29:18.570 --> 00:29:22.080 which is, which was a HIV prevention CRT
- $596~00{:}29{:}22.080 \dashrightarrow 00{:}29{:}25.560$ and 30 Botswana communities that was conducted
- $597\ 00:29:25.560 \longrightarrow 00:29:27.810$ between 2013 and 2018.
- $598~00{:}29{:}27.810 \dashrightarrow 00{:}29{:}32.810$ And this trial was to assess whether an intervention package
- $599\ 00:29:33.090 \longrightarrow 00:29:35.220$ will reduce HIV incidents.
- $600\ 00{:}29{:}35.220 \dashrightarrow 00{:}29{:}38.670$ So in this trial, 15 communities were randomized
- $601\ 00:29:38.670 --> 00:29:41.100$ to receive intervention package
- 602 00:29:41.100 --> 00:29:45.210 that included HIV testing, linkage to care,
- $603\ 00{:}29{:}45.210$ --> $00{:}29{:}48.840$ and early ART initiation for those who are HIV positive
- $604\ 00:29:48.840 \longrightarrow 00:29:50.430$ as well as increased access
- $605\ 00:29:50.430 \longrightarrow 00:29:53.130$ to voluntary medical male circumcision.
- 606 00:29:53.130 --> 00:29:55.350 And the other 15 communities
- $607\ 00:29:55.350 \longrightarrow 00:29:57.250$ were randomized to a standard of care.
- $608\ 00:29:58.290 \longrightarrow 00:30:00.450$ So in the primary analysis they found
- $609\ 00:30:00.450 \longrightarrow 00:30:02.760$ that there were decreased incident rates
- $610\ 00:30:02.760 \longrightarrow 00:30:04.790$ and increased file suppression rates
- $611\ 00:30:04.790 \longrightarrow 00:30:06.090$ in the intervention communities

- $612\ 00{:}30{:}06.090 \dashrightarrow 00{:}30{:}08.340$ compared to control communities.
- $613~00{:}30{:}08.340 \dashrightarrow 00{:}30{:}12.030$ And in our application, we are accounting for non-compliance
- $614\ 00:30:12.030$ --> 00:30:15.150 to the components where we analyzed the individual's
- $615\ 00:30:15.150 \longrightarrow 00:30:16.860$ spillover, total, and over effects
- 616 00:30:16.860 --> 00:30:18.210 of the package intervention
- $617\ 00:30:18.210$ --> 00:30:21.993 that was received on behavioral and clinical outcomes.
- $618\ 00:30:23.340 \longrightarrow 00:30:26.070$ And here, we consider the communities
- $619\ 00:30:26.070 \longrightarrow 00:30:28.680$ to be the misspecified interference sets
- $620\ 00:30:28.680 \longrightarrow 00:30:31.000$ and we determined the true exposures
- 621 00:30:32.817 --> 00:30:35.610 using phylogenetics data, which we consider
- $622\ 00:30:35.610 --> 00:30:37.623$ as our validation data.
- $623~00{:}30{:}40.680 \dashrightarrow 00{:}30{:}44.640$ So the phylogenetic data was obtained from the study shown
- $624\ 00:30:44.640 \longrightarrow 00:30:46.230$ in the beginning of this section
- $625\ 00:30:46.230 \longrightarrow 00:30:49.470$ where they found viral transmission chains
- $626\ 00:30:49.470 \longrightarrow 00:30:52.620$ that crossed multiple clusters.
- $627\ 00{:}30{:}52.620 {\: \hbox{--}}{>}\ 00{:}30{:}56.340$ So here, they approached HIV positive individuals
- $628~00{:}30{:}56.340 {\:{\circ}{\circ}{\circ}}>00{:}30{:}59.340$ in the study and obtained blood samples from them
- $629\ 00{:}30{:}59.340 \dashrightarrow 00{:}31{:}02.940$ and they were able to sequence their viral genomes
- $630\ 00:31:02.940 \longrightarrow 00:31:05.940$ and construct HIV clusters
- 631 00:31:05.940 --> 00:31:08.910 where a participants group in the same viral cluster
- $632\ 00{:}31{:}08.910 \dashrightarrow 00{:}31{:}12.693$ were implied to be from the same viral transmission chain.
- $633~00{:}31{:}13.860 \dashrightarrow 00{:}31{:}17.820$ So here, each viral cluster they found to be composed
- $634\ 00:31:17.820 \longrightarrow 00:31:19.498$ of two to 27 participants who were
- $635\ 00:31:19.498 --> 00:31:21.873$ from one to 16 communities.

- $636\ 00:31:23.190 --> 00:31:26.040$ And there were several considerations that we had to make
- $637~00{:}31{:}27.069 \dashrightarrow 00{:}31{:}31.800$ by using the phylogenetic data as our validation data
- $638\ 00{:}31{:}31.800 \dashrightarrow 00{:}31{:}35.100$ because the viral clusters only captured participants
- $639\ 00:31:35.100 --> 00:31:37.680$ were infected by the same HIV strain.
- $640~00{:}31{:}37.680 \dashrightarrow 00{:}31{:}42.090$ And would not necessarily represent a participant's
- $641\ 00:31:42.090 \longrightarrow 00:31:43.653$ entire true interference set.
- $642\ 00:31:45.360 \longrightarrow 00:31:47.160$ So we had to make some assumptions
- $643\ 00:31:47.160 --> 00:31:50.070$ to obtain the true spillover exposure
- 644 00:31:50.070 --> 00:31:53.430 using this phylogenetic data where the first,
- $645\ 00:31:53.430 \longrightarrow 00:31:56.010$ we consider the connections observed
- $646\ 00:31:56.010 \longrightarrow 00:31:59.127$ within the viral cluster were representative
- $647\ 00:32:00.564 \longrightarrow 00:32:02.567$ of the participants
- $648\ 00:32:02.567 --> 00:32:06.693$ who were HIV positive in ik's true interference set.
- 649 00:32:08.190 --> 00:32:11.600 And then we also assume transportability
- $650\ 00{:}32{:}11.600 \dashrightarrow 00{:}32{:}15.180$ of the measurement error process where we assume
- $651\ 00:32:15.180 \longrightarrow 00:32:16.620$ that the inter-cluster interactions
- $652\ 00:32:16.620 --> 00:32:18.930$ that we observed from the viral clusters
- $653\ 00{:}32{:}18.930 \longrightarrow 00{:}32{:}22.590$ would've been the same among those who were HIV negative.
- $654\ 00:32:22.590 \longrightarrow 00:32:24.600$ And lastly, we considered
- $655~00{:}32{:}24.600 \dashrightarrow 00{:}32{:}26.460$ that because those who are HIV positive
- $656\ 00:32:26.460 --> 00:32:28.200$ might not have the same characteristics
- $657\ 00{:}32{:}28.200 \dashrightarrow 00{:}32{:}31.800$ for intervention uptake as those who are HIV negative.
- $658\ 00:32:31.800 \longrightarrow 00:32:35.400$ We derived the true spillover exposure
- $659\ 00:32:35.400 --> 00:32:39.123$ based on a weighted average of cluster intervention uptake.

- $660\ 00:32:40.170 --> 00:32:44.490$ So for example, if there were five participants observed
- $661\ 00:32:44.490 \longrightarrow 00:32:45.870$ in the viral cluster
- $662\ 00{:}32{:}45.870 \dashrightarrow 00{:}32{:}48.870$ from two different randomization communities,
- $663\ 00:32:48.870 --> 00:32:51.540$ then we take the intervention uptake
- $664\ 00:32:51.540 \longrightarrow 00:32:53.220$ of these two communities
- 665 00:32:53.220 --> 00:32:55.800 and waited by the portion of participants
- $666\ 00{:}32{:}55.800 \dashrightarrow 00{:}32{:}59.583$ from each community that were observed in the viral cluster.
- $667\ 00{:}33{:}03.510 \dashrightarrow 00{:}33{:}08.510$ And so here are some details on the intervention components
- $668\ 00:33:09.300 \longrightarrow 00:33:10.353$ for the study.
- 669 00:33:11.280 --> 00:33:12.690 I don't know, do I have enough to?
- 670 00:33:12.690 --> 00:33:16.260 Okay, so basically,
- $671\ 00{:}33{:}16.260 {\:{--}{>}\:} 00{:}33{:}18.780$ there are four components to this intervention package
- $672\ 00:33:18.780 --> 00:33:21.930$ and these four components were eligible
- $673\ 00:33:21.930 \longrightarrow 00:33:24.480$ to different study populations.
- $674\ 00{:}33{:}24.480 \dashrightarrow 00{:}33{:}27.720$ So here are the eligibility criteria that we had considered
- $675~00{:}33{:}27.720 \dashrightarrow 00{:}33{:}32.610$ for our application where for testing, we considered
- $676\ 00:33:32.610 \longrightarrow 00:33:34.950$ that they were eligible for testing if they did not have
- $677\ 00{:}33{:}34.950 \dashrightarrow 00{:}33{:}38.490$ documented HIV positive status prior to baseline
- $678~00:33:38.490 \longrightarrow 00:33:41.460$ and participants were eligible for HIV care
- $679\ 00{:}33{:}41.460 \dashrightarrow 00{:}33{:}45.633$ and ART initiation if they were HIV positive at baseline.
- $680\ 00:33:46.560 \longrightarrow 00:33:51.060$ And for circumcision, we considered someone
- $681\ 00:33:51.060 \longrightarrow 00:33:53.640$ to be eligible for this treatment
- $682\ 00:33:53.640 --> 00:33:56.190$ if they were an HIV negative male at baseline
- $683\ 00:33:56.190 \longrightarrow 00:33:57.813$ who had not been circumcised.

- $684\ 00:33:59.430 \longrightarrow 00:34:01.680$ And we also considered several definitions
- $685~00{:}34{:}01.680 \dashrightarrow 00{:}34{:}04.920$ of the individual exposure which were receiving
- $686\ 00:34:04.920 \longrightarrow 00:34:07.980$ at least one of these intervention components
- $687\ 00:34:07.980 \longrightarrow 00:34:10.080$ or receiving all eligible components
- $688\ 00:34:10.080 \longrightarrow 00:34:11.733$ versus some or none of them.
- $689\ 00:34:13.345 --> 00:34:16.173$ In this paper, we also considered three outcomes.
- $690\ 00{:}34{:}17.040 \dashrightarrow 00{:}34{:}19.890$ First was a behavioral outcome that we defined
- 691 00:34:19.890 --> 00:34:22.440 as a sexual risk behavior score
- $692\ 00:34:22.440 \longrightarrow 00:34:25.530$ and these is defined as the number
- $693~00{:}34{:}25.530 \dashrightarrow 00{:}34{:}30.377$ of self-reported behaviors that they had reported
- $694\ 00:34:30.377 \longrightarrow 00:34:33.453$ at their survey interview at one year post baseline.
- $695\ 00:34:34.560 \longrightarrow 00:34:36.990$ And then we also looked at two clinical outcomes,
- $696\ 00:34:36.990 \longrightarrow 00:34:39.690$ which were viral load at one year post baseline
- $697\ 00:34:39.690 \longrightarrow 00:34:41.913$ and HIV incidents by the end of the study.
- $698\ 00:34:45.840 \longrightarrow 00:34:47.640$ Before we looked at the effect
- $699\ 00:34:47.640 --> 00:34:50.370$ of receiving the individual components,
- 700~00:34:50.370 --> 00:34:54.300 we first assessed the overall effect of being assigned
- 701 00:34:54.300 --> 00:34:56.820 to an intervention cluster versus control
- $702~00{:}34{:}56.820 \dashrightarrow 00{:}35{:}01.233$ on these three outcomes where the ITT estimates
- $703\ 00:35:01.233 --> 00:35:05.670$ were conducted assuming that the interference sets
- $704\ 00:35:05.670 --> 00:35:06.783$ were communities.
- $705~00{:}35{:}07.890 \dashrightarrow 00{:}35{:}12.890$ And we see that there was a minimal overall effect
- $706\ 00:35:15.060 --> 00:35:20.060$ of cluster assignment on decreasing sexual risk behaviors.

- $707\ 00:35:20.070 \longrightarrow 00:35:23.490$ But there was significant effect on viral load
- $708~00{:}35{:}23.490 \dashrightarrow 00{:}35{:}27.040$ and incidents where our findings echoed the ones
- $709\ 00:35:27.960 --> 00:35:30.450$ from the primary analysis where they found
- $710\ 00:35:30.450 \longrightarrow 00:35:31.620$ increased viral suppression
- $711\ 00:35:31.620 \longrightarrow 00:35:34.080$ and decreased incidents for clusters assigned
- $712\ 00:35:34.080 \longrightarrow 00:35:37.320$ to intervention and versus control.
- $713\ 00:35:37.320 \longrightarrow 00:35:39.000$ And after bias correction we see
- $714\ 00:35:39.000 \longrightarrow 00:35:41.403$ that these effects are again amplified,
- 715 00:35:42.750 --> 00:35:45.360 which was expected due to the high levels
- $716\ 00:35:45.360 \longrightarrow 00:35:48.150$ of inter-cluster mixing where say,
- $717\ 00:35:48.150 --> 00:35:51.540$ some preventative measures from intervention communities
- $718\ 00:35:51.540 \longrightarrow 00:35:54.240$ may have gone into the control communities
- 719 00:35:54.240 --> 00:35:56.670 and some incidents observed in intervention communities
- 720 00:35:56.670 --> 00:35:59.463 may have been attributable to control communities.
- $721\ 00:36:03.027 \longrightarrow 00:36:04.650$ And we also looked at the effect
- 722 00:36:04.650 --> 00:36:06.880 of receiving at least one component
- $723\ 00:36:07.808 \longrightarrow 00:36:12.120$ on essential risk behavior score where we see
- $724\ 00:36:12.120$ --> 00:36:15.030 that after applying our bias correction method,
- $725\ 00:36:15.030 \longrightarrow 00:36:16.620$ that there was a significant total
- $726\ 00:36:16.620 --> 00:36:21.390$ and overall effect of receiving at least one component
- $727\ 00:36:21.390 --> 00:36:23.763$ on decreased sexual risk behaviors.
- 728 00:36:26.880 --> 00:36:30.030 And there was also a significant individual effect
- $729\ 00:36:30.030 \longrightarrow 00:36:32.520$ of receiving both HIV care
- $730~00{:}36{:}32.520 \dashrightarrow 00{:}36{:}37.083$ and ART on decreased viral load which was expected.
- 731 00:36:41.880 --> 00:36:46.789 So here, we proposed methods to bias correct

 $732\ 00:36:46.789$ --> 00:36:50.130 causal effects estimated underspecified interference

733 00:36:50.130 --> 00:36:54.060 sets in a CRT, although our methods are not restricted

 $734\ 00:36:54.060 \longrightarrow 00:36:57.063$ to the setting can be applied to broader settings as well.

 $735\ 00:36:59.010 \longrightarrow 00:37:01.230$ And to use our regression calibration method,

736 00:37:01.230 --> 00:37:03.690 we had to assume that both the measurement error

 $737\ 00:37:03.690 \longrightarrow 00:37:06.003$ and outcome models were correctly specified.

 $738\ 00:37:08.070 \longrightarrow 00:37:09.540$ And we also made some assumptions

 $739\ 00:37:09.540 --> 00:37:11.370$ on the measurement error structure.

740~00:37:11.370 --> 00:37:16.370 So we proposed for a third paper and IPW-based method

 $741\ 00:37:19.290 \longrightarrow 00:37:21.900$ where parametric assumptions on the outcome model

742 00:37:21.900 --> 00:37:23.940 were not required and also we didn't need

 $743\ 00:37:23.940 --> 00:37:26.340$ to make assumptions on the additive

 $744\ 00:37:26.340 \longrightarrow 00:37:29.373$ or non-differential nature of the measurement error process.

 $745\ 00{:}37{:}33.958 {\:\hbox{--}}{>}\ 00{:}37{:}38.958$ Okay, so propensity score based methods are widely used

 $746~00{:}37{:}38.970 \dashrightarrow 00{:}37{:}42.590$ to estimate intervention effects when characteristics

 $747\ 00:37:42.590 \longrightarrow 00:37:46.560$ of the exposed and unexposed participants may be unbalanced,

748 00:37:46.560 --> 00:37:49.860 which may be an observational setting

 $749\ 00:37:49.860 --> 00:37:53.493$ where the exposure is not randomized.

 $750~00{:}37{:}55.050 \dashrightarrow 00{:}37{:}58.440$ And in particular, we're focused on an IPW estimator

 $751\ 00:37:58.440 \longrightarrow 00:38:00.060$ that has been previously extended

 $752\ 00{:}38{:}00.060 {\:{\mbox{--}}\!>\:} 00{:}38{:}03.630$ to estimate causal effects in the setting of interference.

 $753\ 00:38:03.630$ --> 00:38:07.440 And this is typically done assuming the interference sets

 $754\ 00:38:07.440 \longrightarrow 00:38:09.210$ are known and true.

 $755~00{:}38{:}09.210 \dashrightarrow 00{:}38{:}13.600$ And in this paper, we show that when interference sets

 $756\ 00:38:14.580 --> 00:38:17.701$ are mismeasured and spillover exposures are mismeasured

 $757\ 00:38:17.701 --> 00:38:19.250$ as a consequence, there is an error

 $758\ 00:38:19.250 \longrightarrow 00:38:21.090$ in not only the spillover exposure

759 00:38:21.090 --> 00:38:23.433 but also in the propensity score estimates.

 $760\ 00:38:27.660 \longrightarrow 00:38:29.703$ So for notations, here, we have,

761 00:38:30.960 --> 00:38:33.732 we're outside of the network and cluster setting

 $762\ 00:38:33.732 --> 00:38:37.260$ so we have just i from one to end participants.

 $763\ 00{:}38{:}37.260 \dashrightarrow 00{:}38{:}39.840$ Here, the individual exposure status may depend

 $764\ 00:38:39.840 --> 00:38:42.510$ on observed individual covariates.

 $765\ 00:38:42.510 \longrightarrow 00:38:45.120$ And also, here, we assume

 $766\ 00:38:45.120 -> 00:38:47.970$ the pressure interference assumption as in our second paper.

767 00:38:47.970 --> 00:38:50.880 Although this method doesn't require

 $768\ 00:38:50.880 \longrightarrow 00:38:52.590$ the pressure interference assumption.

 $769\ 00{:}38{:}52.590 \dashrightarrow 00{:}38{:}55.350$ We can also make the neighborhood interference consumption

 $770\ 00:38:55.350 \longrightarrow 00:38:58.743$ if we were working in a setting of social networks.

771 $00:39:00.780 \longrightarrow 00:39:04.530$ In this paper, we define a binary spillover exposure,

772 00:39:04.530 \rightarrow 00:39:06.630 although our methods can be generalized

 $773\ 00:39:06.630 \dashrightarrow 00:39:10.200$ to categorical measures of the spillover exposures as well.

 $774\ 00:39:10.200 \longrightarrow 00:39:12.690$ And here we consider an extension

 $775\ 00:39:12.690 \longrightarrow 00:39:14.100$ of the stratify interference

 $776\ 00:39:14.100 --> 00:39:17.970$ that we made in the previous paper where G,

777 00:39:17.970 --> 00:39:20.760 we define by one if the proportion

 $778\ 00:39:20.760 --> 00:39:23.400$ of treated participants in interference set exceeds

779 00:39:23.400 --> 00:39:25.563 a certain pre-specified threshold.

780 00:39:27.150 --> 00:39:30.243 And again, credential outcomes are indexed by A and G.

 $781\ 00:39:31.303 --> 00:39:34.680$ In this paper, we're interested in the individual spillover

 $782\ 00:39:34.680 \longrightarrow 00:39:36.033$ and total effects.

 $783\ 00:39:39.300 \longrightarrow 00:39:41.940$ So this is the IPW estimator

 $784\ 00:39:41.940 --> 00:39:43.900$ for the average potential outcome

 $785\ 00:39:45.300 \longrightarrow 00:39:47.460$ where in the denominator, we have

 $786\ 00:39:47.460 --> 00:39:51.030$ an estimated joint propensity score for the individual

 $787\ 00:39:51.030 \longrightarrow 00:39:52.890$ and spillover exposures.

 $788\ 00:39:52.890 --> 00:39:55.080$ And this can be expressed as the product

789 00:39:55.080 --> 00:39:57.810 of the individual exposure propensity score

 $790\ 00:39:57.810 \longrightarrow 00:40:00.120$ and the spillover exposure propensity score.

 $791\ 00:40:00.120 \longrightarrow 00:40:01.350$ And these can be estimated

792 00:40:01.350 --> 00:40:04.260 using (indistinct) regression models.

793 00:40:04.260 \rightarrow 00:40:07.260 And we can obtain the variance of this estimator

794 00:40:07.260 --> 00:40:10.978 by bootstrap resampling where we can resample

 $795\ 00:40:10.978 \longrightarrow 00:40:14.520$ at the individual level or at the cluster level

 $796\ 00:40:14.520 --> 00:40:16.710$ if we were working in a setting with clusters

 $797\ 00:40:16.710 \longrightarrow 00:40:18.093$ as in our second paper.

798 00:40:19.860 --> 00:40:23.130 And this estimator is consistent, if the models

799~00:40:23.130 --> 00:40:25.833 for the propensity scores are correctly specified.

 $800\ 00:40:29.100 \longrightarrow 00:40:34.100$ So as in the previous cases when interference specified,

 $801\ 00:40:34.980 \longrightarrow 00:40:38.070$ we would observe G star instead of G.

 $802\ 00{:}40{:}38.070 \dashrightarrow 00{:}40{:}42.630$ And if we were to use G star in the IPW estimator,

 $803\ 00:40:42.630 \longrightarrow 00:40:44.970$ we would get a biased estimate

 $804\ 00:40:44.970 \longrightarrow 00:40:48.840$ because the expected value of this estimator is given

 $805~00{:}40{:}48.840 \dashrightarrow 00{:}40{:}52.680$ by the form shown in the bottom here where we see

 $806\ 00:40:52.680 \longrightarrow 00:40:56.610$ that this estimator is only unbiased if the probability

 $807\ 00:40:56.610 \longrightarrow 00:40:59.190$ observing the true exposure equal to G,

 $808\ 00:40:59.190 \longrightarrow 00:41:01.830$ given that the spillover exposure

 $809\ 00:41:01.830 \longrightarrow 00:41:03.810$ is also equal to G is equal to one,

810~00:41:03.810 --> 00:41:06.083 which means that there's no measurement error.

811 00:41:07.386 --> 00:41:12.300 And also from the form of this expectation, we can also see

812 00:41:12.300 --> 00:41:17.040 that the bias can be eliminated if we divide both terms

 $813\ 00:41:17.040 \longrightarrow 00:41:21.000$ on the right-hand side by this measurement error probability

 $814\ 00:41:21.000 \longrightarrow 00:41:23.973$ and then subtracting away the second term.

 $815\ 00:41:25.415 \longrightarrow 00:41:28.980$ Which is the approach that we took.

 $816~00{:}41{:}28.980 \dashrightarrow 00{:}41{:}32.730$ And this was an approach that was first proposed by brown

 $817\ 00{:}41{:}32.730 \dashrightarrow 00{:}41{:}35.970$ and colleagues in the setting without interference.

818 00:41:35.970 --> 00:41:38.220 And here, we extended this estimator

 $819\ 00:41:38.220 \longrightarrow 00:41:39.933$ to the setting of interference.

 $820\ 00{:}41{:}41.640 \dashrightarrow 00{:}41{:}46.230$ So from this bias corrected IPW estimator, we see

 $821\ 00:41:46.230 \longrightarrow 00:41:49.050$ that on the right-hand side in the first term,

822 00:41:49.050 --> 00:41:53.100 we have the IPW estimator that is estimated

 $823\ 00:41:53.100 \longrightarrow 00:41:54.727$ in the main study.

 $824\ 00:41:54.727 --> 00:41:56.880$ We also have an IPW estimator

 $825\ 00:41:56.880 \longrightarrow 00:42:00.180$ that is estimated in the validation study alone.

 $826\ 00:42:00.180 \longrightarrow 00:42:03.630$ And the measurement error probabilities

- $827\ 00:42:03.630 \longrightarrow 00:42:08.630$ are also estimated in the validation study.
- $828\ 00:42:08.730 \longrightarrow 00:42:12.480$ And because here, we are estimating potential outcomes
- $829\ 00:42:12.480 \longrightarrow 00:42:16.740$ in the validation study, we need to assume generalizability
- 830 00:42:16.740 --> 00:42:17.850 of the potential outcome
- $831\ 00:42:17.850 \longrightarrow 00:42:20.640$ and measurement error process in this study
- 832 $00:42:20.640 \longrightarrow 00:42:22.530$ so that the effects that are estimated
- 833 00:42:22.530 --> 00:42:24.630 in the validation study alone
- $834\ 00:42:24.630 \longrightarrow 00:42:27.720$ would be unbiased for the average effect
- 835 00:42:27.720 --> 00:42:29.770 that would be observed in the main study.
- $836\ 00:42:33.060 \longrightarrow 00:42:36.810$ So using these bias corrected IPW estimators,
- $837\ 00:42:36.810 \longrightarrow 00:42:39.420$ we can obtain a bias corrected estimator
- $838\ 00:42:39.420 \longrightarrow 00:42:42.270$ for the causal effect which is given as contrast
- 839 00:42:42.270 --> 00:42:44.880 between potential outcomes estimated
- $840\ 00:42:44.880 \longrightarrow 00:42:47.790$ using the bias corrected IPW estimators.
- 841 00:42:47.790 --> 00:42:50.963 And here, we can write this estimator using
- 842 00:42:54.510 --> 00:42:57.210 with weights, W here.
- $843\ 00:42:57.210 \longrightarrow 00:42:59.760$ Where the weights are meant to minimize the variance
- $844\ 00:42:59.760 --> 00:43:03.150$ of the bias corrected causal effects
- $845\ 00:43:03.150 \longrightarrow 00:43:06.960$ and the weights are given at the bottom here
- $846\ 00{:}43{:}06.960 {\:{\mbox{--}}}\!> 00{:}43{:}10.620$ where the variance of variance terms can also be estimated
- 847 00:43:10.620 --> 00:43:12.650 using bootstrap resampling.
- $848\ 00{:}43{:}16.770 \dashrightarrow 00{:}43{:}20.670$ So while this estimator directly eliminates the bias,
- $849\ 00:43:20.670 \longrightarrow 00:43:21.990$ it does require the outcome
- $850\ 00:43:21.990 \longrightarrow 00:43:25.440$ to be available in the validation study.
- $851\ 00:43:25.440 \longrightarrow 00:43:28.230$ So when this is not available,
- $852\ 00:43:28.230 --> 00:43:30.000$ we propose an alternative estimator
- 853 00:43:30.000 --> 00:43:32.880 that does not impose this requirement

 $854\ 00{:}43{:}32.880 \rightarrow 00{:}43{:}35.550$ where we've extended methods proposed by rule

 $855\ 00:43:35.550 \longrightarrow 00:43:39.540$ and colleagues to the setting of interference.

 $856\ 00{:}43{:}39.540 \dashrightarrow 00{:}43{:}42.600$ And so this is a regression calibration-based approach

 $857\ 00{:}43{:}42.600 \dashrightarrow 00{:}43{:}46.920$ where first, we assume that we have a continuous measure

 $858\ 00:43:46.920 \longrightarrow 00:43:48.453$ of the spillover exposure.

 $859\ 00{:}43{:}49.440 --> 00{:}43{:}53.310$ And we will predict the true continuous spillover exposures

 $860\ 00:43:53.310 \longrightarrow 00:43:55.200$ given the observed ones.

861 00:43:55.200 --> 00:43:57.330 And then under the exposure mapping

 $862\ 00:43:57.330 \longrightarrow 00:44:00.210$ that we had specified previously with the threshold,

 $863\ 00{:}44{:}00.210 \dashrightarrow 00{:}44{:}04.983$ we would dichotomize this proportion.

 $864\ 00{:}44{:}06.870 {\: -->\:} 00{:}44{:}10.440$ And the regression calibration based IPW estimator

 $865\ 00:44:10.440 --> 00:44:14.850$ would use the predicted binary true exposures

 $866\ 00:44:14.850 \longrightarrow 00:44:16.920$ as well as the propensity scores estimated

867 00:44:16.920 --> 00:44:18.930 under these predictive values.

 $868\ 00:44:18.930 \longrightarrow 00:44:22.980$ And we've shown that as in the previous paper,

 $869\ 00:44:22.980 \longrightarrow 00:44:25.440$ that this estimator is only consistent

 $870\ 00:44:25.440 \longrightarrow 00:44:28.120$ if a linear measurement error model fits the data

 $871\ 00:44:31.860 --> 00:44:34.740$ In this paper, we further consider the case

 $872\ 00:44:34.740 \longrightarrow 00:44:37.110$ where we might observe multiple surrogate

 $873\ 00:44:37.110 --> 00:44:39.180$ interference sets in a study.

 $874~00{:}44{:}39.180 \dashrightarrow 00{:}44{:}44.180$ And this was motivated by our illustrative example of BCPP

 $875\ 00{:}44{:}45.240 \dashrightarrow 00{:}44{:}48.180$ where we may consider a surrogate interference set

 $876\ 00:44:48.180 \longrightarrow 00:44:50.460$ defined by a randomization cluster.

 $877\ 00{:}44{:}50.460 \longrightarrow 00{:}44{:}54.420$ And we can also consider a second surrogate interference set

878 00:44:54.420 --> 00:44:56.880 that is defined by household GPS data,

 $879\ 00:44:56.880 \longrightarrow 00:44:59.790$ which is available in the study.

 $880\ 00:44:59.790 \longrightarrow 00:45:03.930$ So when we have multiple surrogate interference sets,

 $881\ 00{:}45{:}03.930 \longrightarrow 00{:}45{:}08.100$ we propose to first apply our bias corrected estimators,

 $882\ 00{:}45{:}08.100 \dashrightarrow 00{:}45{:}12.780$ either the first or the regression calibration-based one

 $883\ 00:45:12.780 \longrightarrow 00:45:15.600$ to each surrogate interference set individually.

884~00:45:15.600 --> 00:45:18.090 And then we will combine these individual estimates

 $885\ 00{:}45{:}18.090 \dashrightarrow 00{:}45{:}21.270$ using a weighted average estimator to reduce the variance

 $886\ 00:45:21.270 \longrightarrow 00:45:22.713$ of the final estimate.

887 00:45:24.690 --> 00:45:29.310 So the weights are given by C in the bottom here

 $888\ 00:45:29.310$ --> 00:45:32.970 where we would estimate the variance variance matrix

 $889\ 00:45:32.970 \longrightarrow 00:45:37.203$ of the individual bias corrected causal effects.

890~00:45:41.640 --> 00:45:45.060 Here, similar to the second paper, we've applied our methods

 $891\ 00:45:45.060 --> 00:45:48.510$ to BCPP where we analyzed the individual's

 $892\ 00:45:48.510 \longrightarrow 00:45:49.860$ spillover total effects

893 00:45:49.860 --> 00:45:52.680 of receiving at least one intervention component

894 00:45:52.680 --> 00:45:56.703 on sexual risk behaviors one year after study enrollment.

 $895\ 00:45:57.630 \longrightarrow 00:46:01.260$ So as a reminder, the components here are HIV testing,

896 00:46:01.260 --> 00:46:05.730 HIV care, ART and circumcision.

 $897\ 00{:}46{:}05.730 \dashrightarrow 00{:}46{:}09.360$ And here, we consider a binary outcome, which we define

 $898~00{:}46{:}09.360 \dashrightarrow 00{:}46{:}13.050$ by one if a participant had reported having engaged

 $899\ 00:46:13.050 \longrightarrow 00:46:16.383$ in at least 30% of the surveyed risk behaviors.

900 00:46:18.150 --> 00:46:19.440 Here, for application,

 $901\ 00:46:19.440 \longrightarrow 00:46:22.050$ we consider the randomization clusters

 $902\ 00{:}46{:}22.050 --> 00{:}46{:}25.680$ or communities as our first surrogate interference set.

 $903~00{:}46{:}25.680$ --> $00{:}46{:}29.580$ And we also consider a second surrogate interference set

 $904\ 00:46:29.580 \longrightarrow 00:46:32.910$ that is defined by smaller geographical plots.

 $905\ 00{:}46{:}32.910 \dashrightarrow 00{:}46{:}37.910$ And in these geographical plots, they comprised

 $906\ 00{:}46{:}37.950 \dashrightarrow 00{:}46{:}41.700$ of participant two to 18 participants on average,

907 00:46:41.700 --> 00:46:44.100 which were much smaller than randomization clusters,

908 00:46:44.100 --> 00:46:46.593 which were about 400 participants each.

909 00:46:48.927 --> 00:46:52.440 And in both of these interference sets,

910 00:46:52.440 --> 00:46:56.660 we define the spillover exposure to be one if at least 25%

911 00:46:56.660 --> 00:46:59.370 of participants in the inference set received

 $912\ 00:46:59.370 \longrightarrow 00:47:02.550$ at least one intervention component.

913 $00:47:02.550 \longrightarrow 00:47:04.230$ And as in the second paper,

 $914\ 00:47:04.230 \longrightarrow 00:47:06.520$ we determine the true spillover exposures

915 00:47:07.967 --> 00:47:09.467 from the phylogenetic dataset.

 $916\ 00:47:12.660 \longrightarrow 00:47:15.600$ So here are the risk differences

917 00:47:15.600 --> 00:47:18.750 of receiving at least one intervention component

 $918\ 00:47:18.750 --> 00:47:20.980$ on self-reported sexual risk behaviors

919 00:47:21.930 --> 00:47:24.750 where we compare the estimates obtained when we consider

920 $00:47:24.750 \longrightarrow 00:47:27.270$ communities at the randomization clusters

921 00:47:27.270 --> 00:47:30.963 or the geographical plot as to the interference sets.

- 922 00:47:32.070 --> 00:47:35.190 And we compare these to the bias corrected estimates
- 923 00:47:35.190 --> 00:47:37.740 where here, I'm presenting the estimates
- 924 00:47:37.740 --> 00:47:39.990 of coming from the weighted average
- $925\ 00{:}47{:}39.990 \dashrightarrow 00{:}47{:}43.080$ of the bias corrected estimates applied individually
- 926 00:47:43.080 --> 00:47:46.293 to the community and to the geographical plots.
- 927 00:47:48.120 --> 00:47:52.181 Where here, under the circuit interference,
- $928\ 00:47:52.181 \longrightarrow 00:47:56.160$ as we see that most effects were null.
- 929 00:47:56.160 --> 00:47:59.610 However, after bias correction, we see
- 930 00:47:59.610 --> 00:48:04.610 that there is a beneficial AIE when G is equal to one
- 931 00:48:04.770 --> 00:48:08.010 and beneficial ASP when A is equal to one.
- $932\ 00:48:08.010 \longrightarrow 00:48:10.740$ Which means that for participants
- $933\ 00:48:10.740$ --> 00:48:13.920 who received at least one component of the intervention,
- 934 00:48:13.920 --> 00:48:16.950 if there were in the presence of at least 25%
- $935\ 00:48:16.950 \longrightarrow 00:48:19.590$ of participants who also received the intervention,
- $936\ 00:48:19.590 \longrightarrow 00:48:22.050$ that they had decreased risk behaviors.
- 937 00:48:22.050 --> 00:48:24.400 And likewise for ASP one,
- 938 00:48:27.660 --> 00:48:30.160 for participants who did receive the intervention,
- 939 00:48:32.430 --> 00:48:37.247 if they were exposed to at least 25% of participants
- $940\ 00:48:37.247 \longrightarrow 00:48:41.280$ in the interference set who also received the intervention,
- 941 00:48:41.280 --> 00:48:44.253 then there risk behavior fears were also reduced.
- 942 00:48:45.090 --> 00:48:46.203 But on the other hand,
- 943 00:48:47.647 --> 00:48:51.480 if a participant did not receive at least one component
- $944\ 00:48:51.480 \longrightarrow 00:48:56.010$ and greater than 75% of those interference

- 945 00:48:56.010 --> 00:48:58.740 that also did not receive the intervention,
- $946\ 00:48:58.740 \longrightarrow 00:49:02.730$ then this had an adverse effect on the risk behaviors.
- 947 00:49:02.730 --> 00:49:07.500 So overall, we see that a participant's risk behaviors
- 948 00:49:07.500 --> 00:49:10.110 are influenced by their own treatment
- 949 00:49:10.110 --> 00:49:14.280 and also in synergy with the treatment received
- $950\ 00:49:14.280 \longrightarrow 00:49:16.773$ by those in their interference set.
- 951 00:49:21.270 --> 00:49:23.820 So to wrap up,
- $952\ 00:49:23.820$ --> 00:49:27.000 so we proposed several bias corrected estimators,
- $953\ 00:49:27.000 --> 00:49:29.220$ which serve to decrease the bias in assessment
- 954 00:49:29.220 --> 00:49:32.100 of causal effects so that future intervention strategies
- $955\ 00{:}49{:}32.100 {\:{\mbox{--}}\!>\:} 00{:}49{:}35.133$ can be more efficiently designed and interpreted.
- $956\ 00:49:36.300 \longrightarrow 00:49:38.400$ And our methods assume
- 957 00:49:38.400 --> 00:49:41.550 that we have suitable validation study that provides us
- $958\ 00:49:41.550 --> 00:49:44.130$ with true measures of the interference set.
- $959\ 00:49:44.130 --> 00:49:46.630$ However, as we see from our application
- $960\ 00:49:47.849 \longrightarrow 00:49:49.560$ that an exposure contamination dataset
- 961 00:49:49.560 --> 00:49:52.980 or a phylogenetic dataset are still imperfect measures
- 962 00:49:52.980 --> 00:49:55.290 of true social connections,
- $963\ 00:49:55.290 \longrightarrow 00:49:56.310$ although we do assume
- $964~00{:}49{:}56.310 \longrightarrow 00{:}49{:}59.430$ that these are more accurate than interference sets defined
- 965 00:49:59.430 --> 00:50:02.010 by general say spatial boundaries
- $966\ 00:50:02.010 \longrightarrow 00:50:03.543$ or administrative boundaries.
- 967 00:50:06.180 --> 00:50:08.430 And we propose for future extensions
- 968 00:50:08.430 --> 00:50:10.780 that we can perform sensitivity analysis
- 969 00:50:12.507 --> 00:50:13.860 on departures from the assumptions

970 00:50:13.860 --> 00:50:15.610 that are made in this dissertation.